THE MAXIMAL VARIATION OF A BOUNDED MARTINGALE

BY

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ABSTRACT

Let $\chi_0^n = \{X_i\}_0^n$ be a martingale such that $0 \le X_i \le 1$; $i = 0, \dots, n$. For $0 \le p \le 1$ denote by \mathcal{M}_p^n the set of all such martingales satisfying also $E(X_0)=p$. The *variation* of a martingale χ_0^n is denoted by $V(\chi_0^n)$ and defined by $V(\chi_0^n)$ = $E(\sum_{i=0}^{n-1} |X_{i+1}-X_i|)$. It is proved that

$$
\lim_{n\to\infty}\left\{\sup_{x_0^n\in\mathcal{M}_p^n}\left[\frac{1}{\sqrt{n}}V(\chi_0^n)\right]\right\}=\phi(p),
$$

where $\phi(p)$ is the well known normal density evaluated at its p-quantile, i.e.

$$
\phi(p) = \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}x_p^2) \text{ where } \int_{-\infty}^{x_p} \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}x^2) dx = p.
$$

A sequence of martingales χ_0^n , $n = 1, 2, \cdots$ is constructed so as to satisfy $\lim_{n \to \infty} (1/\sqrt{n}) V(x_0^n) = \phi(p).$

I. Introduction

For a martingale $\chi_0^n = \{X_i\}_0^n$ we define the variation by $V(\chi_0^n) =$ $E(\sum_{i=0}^{n-1} |X_{i+1}-X_i|)$. We are interested in this variation for bounded martingales, say $0 \le X_i \le 1$, $i = 0, 1, 2, \cdots$. For any $p: 0 \le p \le 1$ denote by \mathcal{M}_p^p the set of all *n*-martingales bounded in [0,1] and satisfying $E(X_0) = p$ ($E(X)$) denotes the expectation of X).

A rather easy consequence of a well known property of martingales and the Cauchy-Schwartz inequality is that

$$
(1.1) \t V(\chi_0^n) \leq \sqrt{p(1-p)} \cdot \sqrt{n}
$$

for every $\chi_0^n \in \mathcal{M}_n^n$. In particular if $\{X_i\}_0^\infty$ is an infinite martingale with $E(X_0) = p$ and χ_0^n is its truncation at stage *n*, then (1.1) holds for $n = 1, 2, \dots$. However this

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is not the strongest statement possible in this case since from the convergence of ${X_i}^{\infty}_{0}$ it can be shown that for any such ∞ -martingale

(1.2)
$$
\lim_{n \to \infty} \frac{1}{\sqrt{n}} V(\chi_0^n) = 0.
$$

The question we are interested in is: Is \sqrt{n} the least upper bound for the order of magnitude of $V(\chi_0^n)$? Since obviously there are *n*-martingales with $V(\chi_0^n)$ of lower order of magnitude, the question is: Is there a function $f(p)$; $f(p) > 0$ for $0 < p < 1$; such that for each $0 \le p \le 1$ and a positive integer *n* there exists an *n*-martingale $\chi_0^n \in \mathcal{M}_p^n$ satisfying

$$
(1.3) \t\t V(\chi_0^n) \ge f(p) \sqrt{n} \t?
$$

Notice that in view of (1.2) it is impossible to satisfy (1.3) with the χ_0 ⁿ being the truncations of the same ∞ -martingale. An affirmative answer to the above stated question would imply: There exists $f(p)$: $f(0) = f(1) = 0$ and $f(p) > 0$ for $0 < p < 1$ such that

(1.4)
$$
f(p) \leq \sup_{\chi_0^n \in \mathcal{M}_p^n} \left[\frac{1}{\sqrt{n}} V(\chi_0^n) \right] \leq \sqrt{p(1-p)}.
$$

It turns out that a result much stronger than (1.4) can be achieved, namely

(1.5)
$$
\lim_{n \to \infty} \left\{ \sup_{\chi_0^n \in \mathcal{M}_p^n} \left[\frac{1}{\sqrt{n}} V(\chi_0^n) \right] \right\} = \phi(p)
$$

where

$$
\phi(p) = \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}x_p^2); \qquad \int_{-\infty}^{x_p} \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}x^2) \, dx = p.
$$

Thus, not only is $\text{Sup}_{x_0 \in \mathcal{M}_n^n} [(1/\sqrt{n}) V(x_0)]$ bounded away from 0 but it is a converging sequence, the limit of which is, amazingly enough, the well known normal density function evaluated at its p-quantile.

A by-product of the proof of this result is a construction of a sequence of martingales $\chi_0^n \in \mathcal{M}_p^n$; $n = 1, 2, \cdots$ for which

$$
\lim_{n\to\infty}\left[\frac{1}{\sqrt{n}}V(\chi_0^n)\right]=\phi(p).
$$

Our interest in the variation of bounded martingales came up through game theory. It turns out that the speed of convergence of the values of certain repeated games with incomplete information is given by $(1/n)V(y_0^n)$ for some $\chi_0^n \in \mathcal{M}_n^n$. Using the results of this paper one can find a sequence of such games with asymptotic value 0, for which the values $v_n(p)$ satisfy $\lim_{n\to\infty} \sqrt{nv_n(p)}$ $\phi(p)$. These are in some sense repeated games with the slowest rate possible for releasing information (see [3]).

We are indebted to David Gilat for calling our attention to an inaccuracy in the first version of the paper.

2. Preliminary results and statement of the main theorem

We are considering *n*-martingales and ∞ -martingales bounded in [0, 1] with a given and fixed expectation p, i.e., $\{X_i\}_0^n$ and $\{X_i\}_0^\infty$ such that $0 \le X_i \le 1$ Vi and $E(X_0) = p$; $0 \le p \le 1$ (hence also $E(X_i) = p \forall i$). We denote such *n*-martingales by χ_0^* (or χ_0^*) and the set of all such martingales by \mathcal{M}_p^* (or \mathcal{M}_p^*).

DEFINITION 2.1 The *n*-stage variation of a martingale χ_0^{∞} (or χ_0^{∞} for $m \ge n$) is denoted by $V(\chi_0^n)$ and defined by

$$
V(\chi_0^n) = E\bigg(\sum_{i=0}^{n-1} |X_{i+1} - X_i|\bigg).
$$

The following two Theorems may be partially or fully known. However, we state and prove them here for the sake of completeness and mainly to clarify the significance of our main result (Theorem 2.5).

THEOREM 2.2 For all $p, 0 \leq p \leq 1$ and $n = 1, 2, \cdots$

(2.1)
$$
\frac{1}{\sqrt{n}}V(\chi_0^n) \leq \sqrt{p(1-p)}
$$

holds for all $\chi_0^* \in \mathcal{M}_p^n$.

PROOF. We recall that since martingale differences are uncorrelated (see e.g. [1]), it follows that

(2.2)
$$
E\left(\sum_{0}^{n-1} (X_{i+1} - X_i)^2\right) = E(X_n - X_0)^2.
$$

Since in our case the martingale is bounded in [0, 1] and $E(X_0) = p$, it is easily seen that the maximal value for $E(X_n - X_0)^2$ is attained when $X_0 \equiv p$ and the distribution of X_n is $Pr(X_n = 1) = p$; $Pr(X_n = 0) = 1 - p$ in which case we have

$$
E(X_n - X_0)^2 = E(X_n^2) - p^2 = p - p^2 = p(1 - p).
$$

Hence, by (2.2)

(2.3)
$$
E\left(\sum_{i=0}^{n-1} (X_{i+1}-x_i)^2\right) \leq p(1-p).
$$

Now we make use of the Cauchy-Schwartz inequality and get (using (2.3))

$$
V(\chi_0^n) = E\left(\sum_{i=0}^{n-1} |X_{i+1} - X_i| \right) \leq \left[E\left(\sum_{i=0}^{n-1} (X_{i+1} - X_i)^2 \right) E\left(\sum_{i=0}^{n-1} 1^2 \right) \right]^{1/2}
$$

$$
\leq \sqrt{p(1-p)} \cdot \sqrt{n},
$$

which completes the proof.

COROLLARY 2.3 For any infinite martingale χ_0^* bounded in [0,1] with $E(X_0) = p$ the *n*-stage variation is bounded by $\sqrt{p(1-p)} \cdot \sqrt{n}$ and consequently $0 \leq \lim_{n \to \infty} \sup V(\chi_0^n)/\sqrt{n} \leq \sqrt{p(1-p)}.$

However, as far as infinite martingales are concerned a stronger result can be obtained, namely

THEOREM 2.4. *For any* ∞ -martingale χ_0^* with $0 \leq X_i \leq 1$; $i = 1, 2, \dots$,

$$
\lim_{n\to\infty}\frac{1}{\sqrt{n}}\,V(\chi_0^n)=0\,.
$$

Before proving this theorem let us notice that there is no hope to strengthen it so much as to prove that $V(\chi_0^*)<\infty$. In fact the following example communicated to us by David Gilat shows that a bounded martingale may have an infinite variation.

EXAMPLE. Perform a symmetric random walk $(p = \frac{1}{2})$ with $X_0 = \frac{1}{2}$ and a step size $\frac{1}{8}$. Reduce the step size to $\frac{1}{32}$ as soon as you reach one of the points $\frac{1}{2} \pm \frac{1}{4}$. In general if the step size was last time reduced at point y to size $\varepsilon_k = 2^{-(2k+1)}$, then reduce it to $\varepsilon_{k+1} = 2^{-(2k+3)}$ as soon as you reach one of the points $y \pm n_k \varepsilon_k$ where $n_k = 2^k$. Doing that for $k = 1, 2, \dots$ we obtain a martingale χ_0^* bounded in [0, 1] (since $\sum_{k=1}^{\infty} n_k \varepsilon_k = \frac{1}{2}$). If we denote by N_k the number of steps of size ε_k made, then clearly

$$
V(\chi_0^*)=\sum_{k=1}^{\infty}\,\varepsilon_kE(N_k).
$$

But $E(N_k) = n_k^2$ (expected duration of a classical ruin game, see e.g. [2] pp. 348), hence,

$$
V(\chi_0^*) = \sum_{k=1}^{\infty} \varepsilon_k n_k^2 = \sum_{k=1}^{\infty} 2^{-(2k+1)} 2^{2k} = \infty.
$$

PROOF OF THEOREM 2.4. Let $\chi_0^* = \{X_n\}_0^*$ be an ∞ -martingale bounded in [0, 1]. By (2.2), χ_0^* is (uniformly) bounded in L_2 , hence χ_0^* converges in L_2 to a random variable say X_{∞} . So $\forall \varepsilon > 0$, $\exists k$ s.t. $||X_n - X_{k+n}||_2 < \varepsilon$ for all *n*. Now

$$
\frac{1}{\sqrt{k+n}} V(\chi_0^{k+n}) = \frac{1}{\sqrt{k+n}} \sum_{i=0}^{k+n-1} E(|X_{i+1} - X_i|)
$$

=
$$
\frac{1}{\sqrt{k+n}} \left[\sum_{i=0}^{k-1} E(|X_{i+1} - X_i|) + \sum_{i=k}^{k+n-1} E(|X_{i+1} - X_i|) \right].
$$

Using the Cauchy-Schwartz inequality for each of the two sums and then applying (2.2) we get

$$
\frac{1}{\sqrt{k+n}}V(\chi_0^{k+n}) \leq \frac{\sqrt{k}}{\sqrt{k+n}} \left[\sum_{i=0}^{k-1} E(X_{i+1} - X_i)^2 \right]^{1/2} \n+ \frac{\sqrt{n}}{\sqrt{k+n}} \left[\sum_{i=k}^{k+n-1} E(X_{i+1} - X_i)^2 \right]^{1/2} \n\leq \frac{\sqrt{k}}{\sqrt{k+n}} \|X_k - X_0\|_2 + \frac{\sqrt{n}}{\sqrt{k+n}} \|X_{k+n} - X_k\|_2.
$$

Letting $n \to \infty$ we conclude that $\forall \varepsilon > 0 \exists k$ s.t

$$
\lim_{n\to\infty}\frac{1}{\sqrt{k+n}}V(\chi_0^{k+n})<\varepsilon,
$$

in other words

$$
\lim_{n\to\infty}\frac{1}{\sqrt{n}}\,V(\chi_0^n)=0
$$

as claimed. We now turn to the main result of the paper:

THEOREM 2.5 (The Main Theorem). For any p ; $0 \le p \le 1$:

(2.4)
$$
\lim_{n\to\infty}\left\{\sup_{\chi_0^n\in\mathcal{M}_p^n}\left[\frac{1}{\sqrt{n}}V(\chi_0^n)\right]\right\}=\phi(p),
$$

where $\phi(p)$ is the standard normal density function evaluated at its p-quantile (i.e. $\phi(p) = (1/\sqrt{2\pi})\exp(-\frac{1}{2}x_p^2)$ where $\int_{-\infty}^{x_p} (1/\sqrt{2\pi})\exp(-\frac{1}{2}x^2) dx = p$.

In view of the length and the technical complication of the proof we provide in the next section the heuristic arguments that lead, and in fact have led us, to the result. The formal proof is given in section 4.

3~ **The heuristics of the result**

Let

$$
\phi_n(p) = \sup_{\chi_0^n \in \mathcal{M}_p^n} \left[\frac{1}{\sqrt{n}} V(\chi_0^n) \right].
$$

Assuming $X_0 = p$, $\phi_n(p)$ clearly satisfies the following recursion formula:

$$
\sqrt{n+1} \, \phi_{n+1}(p) = \max_{\{X_i | E(X_i) = p\}} \{ E(|X_i - p|) + \sqrt{n} E(\phi_n(X_i)) \}
$$

(3.1)
\n
$$
= \max_{\{X_i \mid E(X_i) = p\}} \{P(X_i \ge p) \times [E(X_i - p | X_i \ge p) + \sqrt{n}E(\phi_n(X_i) | X_i \ge p)] + P(X_i < p)[E(p - X_i | X_i < p) + \sqrt{n}E(\phi_n(X_i) | X_i < p)]\}.
$$

Assuming that ϕ_n is concave (an assumption based on observation of ϕ_1 and ϕ_2), then for any X_1 , the expression to be maximized in (3.1) is increased if X_1 is replaced by \tilde{X}_1 which assumes only the two values $p + \xi$ and $p - \eta$, where $\xi = E(X_1 - p | X_1 \geq p)$ and $\eta = E(p - X_1 | X_1 < p)$. Since in addition $E(X_1) = p$ we replace (3.1) by

$$
(3.2) \ \sqrt{n+1} \, \phi_{n+1}(p) = \max_{(\xi,\eta)\in S(p)} \left\{ \sqrt{n} \left[\frac{\eta}{\xi+\eta} \, \phi_n(p+\xi) + \frac{\xi}{\xi+\eta} \, \phi_n(p-\eta) \right] + \frac{2\xi\eta}{\xi+\eta} \right\}
$$

where $S(p) = \{(\xi, \eta) | 0 \le \xi \le 1 - p; 0 \le \eta \le p\}.$

By concavity of ϕ_n , the expression in the square brackets decreases both in ξ and in η . Since this expression is multiplied by \sqrt{n} , it follows that the points $({\xi_n}, \eta_n)$ at which (3.2) achieves its maximum, satisfy $({\xi_n}, \eta_n) \longrightarrow (0, 0)$. Expanding $\phi_n(p + \xi)$ and $\phi_n(p - \eta)$ we obtain the following approximation:

$$
(3.3) \qquad \sqrt{n+1} \, \phi_{n+1}(p) \cong \{ \sqrt{n} [\phi_n(p) + \xi_n \eta_n \phi''_n(p)] + 2 \xi_n \eta_n / (\xi_n + \eta_n) \}
$$

For any fixed $\xi_n \eta_n$, the expression at the right hand side of (3.3) is maximized when $({\xi_n} + \eta_n)$ is minimized, which is at ${\xi_n} = \eta_n$. We conclude that as a first approximation, the maximum in (3.2) is achieved for $\zeta = \eta$. Motivated by this observation we restrict the domain of maximization $S(p)$ in (3.2) to $S(p) \cap \{\xi = \xi\}$ $\{\eta\} = \{x \mid 0 \le x \le p^*\},\$ where $p^* = \min(p, 1 - p)$. The recursion equation (3.2) is thus replaced by

(3.4)
$$
\sqrt{n+1} \phi_{n+1}(p) = \max_{0 \le x \le p^*} \left\{ \frac{1}{2} \sqrt{n} [\phi_n(p+x) + \phi_n(p-x)] + x \right\}.
$$

Assuming now that $\phi_n(p)$ converge to some function $\varphi(p)$ one gets from (3.4), letting $x = \alpha_n / \sqrt{n}$,

$$
\sqrt{1+\frac{1}{n}}\varphi(p) \cong \operatorname{Max}_{\alpha_n} \left\{ \frac{\alpha_n}{n} + \frac{1}{2} [\varphi(p + \alpha_n/\sqrt{n}) + \varphi(p - \alpha_n/\sqrt{n})] \right\}
$$

$$
\cong \operatorname{Max}_{\alpha_n} \left\{ \frac{\alpha_n}{n} + \varphi(p) + \frac{\alpha_n^2}{2n} \varphi''(p) \right\}
$$

$$
\cong \varphi + \frac{1}{n} \operatorname{Max}_{\alpha_n} \left\{ \alpha_n + \frac{\alpha_n^2}{2} \varphi'' \right\} = \varphi - \frac{1}{2n\varphi''}.
$$

On the other hand $\sqrt{1 + 1/n} \varphi \cong \varphi + (1/2n) \varphi$, thus $\varphi = -1/\varphi''$. In other words we are led to the differential equation

$$
\varphi\varphi''+1=0.
$$

variable z : To solve (3.5) we rewrite it as $-\varphi'(p) = \int_{1/2}^p (1/\varphi) dp$, where we have introduced an integration constant so as to have $\varphi'(\frac{1}{2}) = 0$ which is implied by the symmetry of $\varphi(p)$ around $p = \frac{1}{2}$. Now let $z(p) = -\varphi'(p) = \int_{1/2}^{p} (1/\varphi) dp$, then $z'(p) = 1/\varphi$ and thus $\varphi = dp/dz$. Now replace in (3.5) the variable p by the

$$
\varphi'_z = \varphi'_p \frac{dp}{dz} = \varphi'_p \varphi = - z \varphi,
$$

which implies $\ln \varphi = K - \frac{1}{2}z^2$ or

(3.6)
$$
\varphi = A \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}z^2)
$$

(where K and A are constants). Since $\varphi = dp/dz$ we get

(3.7)
$$
p = c + \int_{-\infty}^{z(p)} A \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}x^2) dx.
$$

Denoting by $F(x)$ the cumulative standard normal distribution we have therefore

$$
\varphi(z) = AF'(z(p)),
$$

$$
(3.9) \t\t\t p = c + AF(z(p)).
$$

Now $\varphi \ge 0$ and $\varphi \ne 0$ implies $A > 0$ from which it follows by (3.7) that $z(p)$ is monotonously increasing with p. Since $\varphi(0) = \varphi(1) = 0$ we have by (3.6) that $z(0) = -\infty$, $z(1) = +\infty$. From (3.7) we thus have

$$
(3.10) \t1 = c + A.
$$

From $\varphi'(\frac{1}{2}) = 0$ we get $z(\frac{1}{2}) = -\varphi'(\frac{1}{2}) = 0$, hence from (3.7)

$$
\frac{1}{2} = c + \frac{1}{2}A.
$$

We conclude from (3.10) and (3.11) that $c = 0$ and $A = 1$, thus finally $\varphi(p) =$ $F'(z)$, $p = F(z)$, i.e. the limit $\varphi(p)$ is the standard normal density evaluated at its p-quantile.

4. Proof of the Main Theorem

First let us introduce the convention of writing p' instead of $1-p$ for $0 \leq p \leq 1$. (Although we will use the prime also for derivative it will be clear from the context which operation it stands for.) Next, for $0 \le p \le 1$ let

$$
S(p) = \{(\xi, \eta)\}_{0 \leq \eta \leq p}^{\infty \leq \epsilon \leq p'}\}
$$

and define two sequences of functions on [0, 1], { φ_n } and { ψ_n } by $\varphi_0 = \psi_0 = 0$ and for $n = 0, 1, 2, \cdots$

$$
(4.1)\ \sqrt{n+1}\ \varphi_{n+1}(p)=\max_{(\xi,\eta)\in S(p)}\bigg[\sqrt{n}\frac{\eta}{\xi+\eta}\ \psi_n(p+\xi)+\sqrt{n}\frac{\xi}{\xi+\eta}\ \psi_n(p-\eta)+\frac{2\xi\eta}{\xi+\eta}\bigg],
$$

$$
\psi_{n+1} = \text{Cav }\varphi_{n+1}.
$$

In (4.1) the expression in the square brackets is defined to be $\psi_n(p)$ when $\xi = \eta = 0$. In (4.2) Cav is the operator of concavification of a function on [0, 1], (i.e. Cav f is the smallest concave function g satisfying $g(p) \geq f(p)$; $0 \leq p \leq 1$).

We first observe that

(4.3)
$$
\varphi_n(0) = \psi_n(0) = \varphi_n(1) = \psi_n(1) = 0 \text{ for } n = 0, 1, 2, \cdots.
$$

It also follows immediately from the definitions that

(4.4)
$$
\varphi_1(p) = \psi_1(p) = 2p(1-p)
$$
 for $0 \le p \le 1$.

LEMMA 4.1. *Every* φ_n *and every* ψ_n *, n* = 0, 1, 2, \cdots *is symmetric around p* = $\frac{1}{2}$.

This is easily proved by induction using (4.1) and observing that the operation Cav conserves the symmetry around $p = \frac{1}{2}$.

LEMMA 4.2.

(4.5)
$$
\lim_{\substack{\lambda \to \in \mathcal{M}_b}} V(\chi_0^1) = \psi_1(p) = 2p(1-p).
$$

PROOF. For any martingale $\chi_0^n \in \mathcal{M}_p^n$ we denote by $V(\chi_0^n | X_0 = \alpha)$ the conditional *n*-stage variation given the value α of X_0 . We shall first show

(4.6)
$$
\sup_{X_1} V(\chi_0^1 | X_0 = \alpha) = 2\alpha (1 - \alpha),
$$

where the Sup is being taken over all r.v. (random variables) in $[0, 1]$ s.t. $E(X_1) = \alpha$. To prove (4.6) take any such X_1 and let $\Omega_1 = \{X_1 > \alpha\}, \Omega_2 = \{X_1 \leq \alpha\},\$ then

$$
V(\chi_0^1 | X_0 = \alpha) = E(|X_1 - \alpha|)
$$

= $P(\Omega_1)E((X_1 - \alpha)|\Omega_1) + P(\Omega_2)E((\alpha - X_1)|\Omega_2)$
= $P(\Omega_1) \cdot \xi + (1 - P(\Omega_1)) \cdot \eta$

where $\xi = E((X_1 - \alpha)|\Omega_1)$ and $\eta = E((\alpha - X_1)|\Omega_2)$. Now

$$
\alpha = E(X_1) = P(\Omega_1)E(X_1|\Omega_1) + (1 - P(\Omega_1))E(X_1|\Omega_2)
$$

= $P(\Omega_1)(\alpha + \xi) + (1 - P(\Omega_1))(\alpha - \eta)$,

which implies $P(\Omega_1) = \eta/(\xi + \eta)$ and hence $V(\chi_0 | X_0 = \alpha) = 2\xi \eta/(\xi + \eta)$. Therefore

(4.7)
$$
\text{Sup } V(\chi_0^1 | X_0 = \alpha) = \sup_{(\xi, \eta) \in S(\alpha)} 2\xi \eta / (\xi + \eta).
$$

In (4.7) the sup is achieved for $\zeta = \alpha' = 1 - \alpha$ and $\eta = \alpha$, which establishes (4.6).

Now if we denote by E_{x_0} the expectation with respect to the r.v. X_0 we have

$$
\begin{aligned} \text{Sup } \; V(\chi_0^1) &= \sup_{(X_0|E(X_0)=p)} E_{X_0} \bigg[\text{Sup } V(\chi_0^1 \,|\, X_0) \bigg] \\ &= \sup_{(X_0|E(X_0)=p)} E \big(2X_0 (1-X_0) \big). \end{aligned}
$$

Finally since $2X_0(1-X_0)$ is a concave function w.r.t. X_0 , the Sup_{$\{x_0|E(X_0)=p\}$} is achieved for $X_0 = p$ which concludes the proof of Lemma 4.2.

THEOREM 4.3. *For* $n = 0, 1, 2, \cdots$ and $0 \leq p \leq 1$

(4.8)
$$
\underset{\chi_0^n \in \mathcal{M}_p^n}{\sup} \left[\frac{1}{\sqrt{n}} V(\chi_0^n) \right] \leq \psi_n(p).
$$

PROOF. By induction on n: For $n = 1$, (4.8) follows from Lemma 4.2. Assume (4.8) is true for $n \le m - 1$ and let us prove it for $n = m$. Since $\psi_m = \text{Cav } \varphi_m$ it is clearly enough to prove that for $0 \leq p \leq 1$

(4.9)
$$
\underset{\underset{\chi_0 \equiv \mu_p}{\chi_0 \in \mu_p}}{\sup} V\left[\frac{1}{\sqrt{m}}V(\chi_0^m)\right] \leq \varphi_m(p).
$$

To prove (4.9) let $\Omega_1 = \{x_1 > p\}$, $\Pi = P(\Omega_1)$ (hence $1 - \Pi = P(\Omega_2)$). We have for any $\{X_i\}_{0}^{m}$, $X_0 \equiv p$

$$
V(\chi_0^m) = \sum_{i=1}^{m-1} E(|X_i - X_{i-1}|)
$$

= $E(|X_0 - p|) + \prod_{i=2}^{m-1} E(|X_i - X_{i-1}| \Omega_i) + (1 - \prod_{i=2}^{m-1} E(|X_i - X_{i-1}| \Omega_2))$

(by induction hypothesis)

$$
\leq E(|X_1 - p|) + \sqrt{m-1}[\Pi \psi_{m-1}(E(X_1 | \Omega_1)) + (1 - \Pi)\psi_{m-1}(E(X_1 | \Omega_2))]
$$

= $\Pi(E(X_1 | \Omega_1) - p) + (1 - \Pi)(p - E(X_1 | \Omega_2))$
+ $\sqrt{m-1}[\Pi \psi_{m-1}(E(X_1 | \Omega_1)) + (1 - \Pi)\psi_{m-1}(E(X_1 | \Omega_2))].$

Let $p+\xi=E(X_1|\Omega_1)$ and $p-\eta=E(X_1|\Omega_2)$, then $\xi\geq 0$, $\eta\geq 0$ and $\Pi \xi - (1 - \Pi)\eta = 0$ which implies $\Pi = \eta / (\xi + \eta)$. From our last inequality we thus obtain

$$
V(\chi_0^m)\leq \sqrt{m-1}\left[\frac{\eta}{\xi+\eta}\psi_{m-1}(p+\xi)+\frac{\xi}{\xi+\eta}\psi_{m-1}(p-\eta)\right]+\frac{2\xi\eta}{\xi+\eta},
$$

and by definition of $\varphi_m(p)$: $V(\chi_0^m) \leq \sqrt{m} \varphi_m(p)$, which concludes the proof of Theorem 4.3.

The following two Lemmas provide bounds for the convexity of the function $\phi(p)$.

LEMMA 4.4. *There exists a constant* $c > 0$ such that

$$
(4.10) \quad \frac{1}{\sqrt{n+1}} \max_{0 \le x \le \min(p,p')} \left[\frac{\sqrt{n}}{2} (\phi(p+x) + \phi(p-x)) + x \right] \ge \phi(p) - c/n^2
$$

for all $0 \leq p \leq 1$; $n = 1, 2, \cdots$.

LEMMA 4.5. *There exists a constant* $K > 0$ *s.t. for* $0 \le p \le 1$

$$
(4.11) \quad \frac{1}{\sqrt{n+1}} \max_{(\xi,\eta)\in S(p)} \left[\sqrt{n} \frac{\eta}{\xi+\eta} \phi(p+\xi) + \sqrt{n} \frac{\xi}{\xi+\eta} \phi(p-\eta) + \frac{2\xi\eta}{\xi+\eta} \right] \leq \phi(p) + K/n^2.
$$

Unfortunately the proofs of these technical Lemmas about $\phi(p)$ are rather **lengthy. They can be found in the Appendix. Proceeding in the proof of our main theorem we define now a new sequence** $\{\tilde{\phi}_n\}^{\infty}$ **of functions on [0, 1] by** $\tilde{\phi}_0 = 0$ and

$$
(4.12)\sqrt{n+1}\,\tilde{\phi}_{n+1}(p)=\max_{(\xi,\eta)\in S(p)}\bigg[\sqrt{n}\frac{\eta}{\xi+\eta}\,\tilde{\phi}_n(p+\xi)+\sqrt{n}\frac{\xi}{\xi+\eta}\,\tilde{\phi}_n(p-\eta)+\frac{2\xi\eta}{\xi+\eta}\bigg]
$$

(Here again the expression in the brackets is defined to be $\bar{\phi}_n(p)$ if $\xi = \eta = 0$.)

PROPOSITION 4.6.

$$
(4.13) \qquad \psi_n(p) \geq \bar{\phi}_n(p) \quad \text{for} \quad 0 \leq p \leq 1; \quad n = 0, 1, 2, \cdots
$$

This follows readily from the definitions (4.1) (4.2) and (4.12).

LEMMA 4.7. *For* $n = 1, 2, \cdots$ *and* $0 \leq p \leq 1$

(4.14)
$$
\underset{\mathbf{x} \delta \in \mathcal{M}_p^n}{\sup} \left[\frac{1}{\sqrt{n}} V(\mathbf{\chi}_0^n) \right] \geq \tilde{\phi}_n(p).
$$

PROOF. For each $n = 1, 2, \cdots$ let us construct for each $p, 0 \leq p \leq 1$ a martingale $\chi_0^n(p) \in \mathcal{M}_p^n$ with variation exactly $\tilde{\phi}_n(p)$, i.e.

$$
(4.15) \qquad \qquad \frac{1}{\sqrt{n}}\,V(\chi_0^n(p))=\tilde{\phi}_n(p).
$$

We do the construction inductively on *n*: For $n = 1$ let $X_0(p) \equiv p$ and $Pr{X_1(p) = 0} = p'$; $Pr{X_1(p) = 1} = p$ then $V(\chi_0^1(p)) = 2pp' = \tilde{\phi}_1(p)$.

Assume now that for n and for each $p, 0 \leq p \leq 1$ there is a martingale $\{X_i(p)\}_0^n$ satisfying (4.15). Let $({\xi_n}, {\eta_n})$ be the point at which the maximum in (4.12) is attained. Define the martingale $\{Z_{\iota}(p)\}_{{\scriptscriptstyle{0}}}^{n+1}$ by

(4.16)

$$
\begin{cases}\nZ_0(p) \equiv p; \quad \Pr\{Z_1(p) = p + \xi_n\} = \eta_n/(\xi_n + \eta_n); \\
\Pr\{Z_1(p) = p - \eta_n\} = \xi_n/(\xi_n + \eta_n), \\
\{Z_1(p) | Z_1(p) = p + \xi_n\}_2^n = \{X_1(p + \xi_n)\}_1^n, \\
\{Z_1(p) | Z_1(p) = p - \eta_n\}_2^n = \{X_1(p - \eta_n)\}_1^n.\n\end{cases}
$$

It follows by (4.12) and (4.16) that

$$
\frac{1}{\sqrt{n+1}} V(Z_0^{n+1}(p)) = \frac{1}{\sqrt{n+1}} \left[\frac{2\xi_n \eta_n}{\xi_n + \eta_n} + \frac{\eta_n}{\xi_n + \eta_n} V(\chi_0^n(p + \xi_n)) + \frac{\xi_n}{\xi_n + \eta_n} V(\chi_0^n(p - \eta_n)) \right]
$$

$$
= \frac{1}{\sqrt{n+1}} \left[\frac{2\xi_n \eta_n}{\xi_n + \eta_n} + \frac{\eta_n}{\xi_n + \eta_n} \sqrt{n} \phi_n(p + \xi_n) + \frac{\xi_n}{\xi_n + \eta_n} \sqrt{n} \phi_n(p - \eta_n) \right]
$$

$$
= \tilde{\phi}_{n+1}(p).
$$

This completes the proof of Lemma 4.7.

LEMMA 4.8. *For* $0 \le p \le 1$; $n = 1, 2, \cdots$

$$
(4.17) \qquad \qquad \tilde{\phi}_n(p) \geq \phi(p) - \frac{\alpha}{\sqrt{n}},
$$

for some constant $\alpha > 0$ *.*

PROOF. We first prove by induction on k that for any $n \ge 1$

(4.18)
$$
\tilde{\phi}_{n+k}(p) \ge \phi(p) - \frac{1}{\sqrt{n+k}} \left[\frac{\sqrt{n}}{2} + \sum_{i=n}^{n+k} \frac{4c}{i \sqrt{i}} \right],
$$

for $k = 0, 1, 2, \cdots$ where c is a constant satisfying (4.10). In fact for $k = 0$, $\tilde{\phi}_n(p) \ge 0 \ge \phi(p)-\frac{1}{2}$. We notice that in (4.10) we may replace the last term $-c/n^2$ by $-4c/(n+1)^2$. Assume now that (4.18) holds for k, then by (4.10)

$$
\tilde{\phi}_{n+k+1}(p) = \frac{1}{\sqrt{n+k+1}} \max_{(\epsilon,\eta)\in S(p)} \left[\sqrt{n+k} \frac{\xi}{\xi+\eta} \tilde{\phi}_{n+k}(p+\xi) + \frac{2\xi\eta}{\xi+\eta} \tilde{\phi}_{n+k}(p-\eta) + \frac{2\xi\eta}{\xi+\eta} \right]
$$
\n
$$
\geq \frac{1}{\sqrt{n+k+1}} \max_{0 \leq x \leq p \land p'} \left\{ \frac{\sqrt{n+k}}{2} (\tilde{\phi}_{n+k}(p+x) + \tilde{\phi}_{n+k}(p-x)) + x \right\}
$$
\n
$$
\geq \frac{1}{\sqrt{n+k+1}} \max_{0 \leq x \leq p \land p'} \left\{ \frac{\sqrt{n+k}}{2} \left[\phi(p+x) + \phi(p-x) + x \right] - \frac{2}{\sqrt{n+k}} \left(\frac{\sqrt{n}}{2} + \sum_{i=n}^{n+k} \frac{4c}{i \sqrt{i}} \right) \right]
$$

$$
\geq \phi(p) - \frac{4c}{(n+k+1)^2} - \frac{1}{\sqrt{n+k+1}} \left(\frac{\sqrt{n}}{2} + \sum_{i=n}^{n+k} \frac{4c}{i \sqrt{i}} \right)
$$

$$
= \phi(p) - \frac{1}{\sqrt{n+k+1}} \left(\frac{\sqrt{n}}{2} + \sum_{i=n}^{n+k+1} \frac{4c}{i \sqrt{i}} \right).
$$

Now by (4.18) for $n = 1$

$$
\tilde{\phi}_{k+1}(p) \geqq \phi(p) - \frac{1}{\sqrt{k+1}} \left(\frac{1}{2} + \sum_{i=1}^{\infty} \frac{4c}{i \sqrt{i}} \right) = \phi(p) - \frac{\alpha}{\sqrt{k+1}}.
$$

Since this holds for $k = 0, 1, \dots$, the proof of the Lemma is completed.

LEMMA 4.9. *For* $0 \leq p \leq 1$ *and* $n = 1, 2, \cdots$

(4.18) 0,(P) --< 4'(P) +/3/\/n

for some constant $\beta > 0$.

PROOF. The proof is almost the same as that of Lemma 4.8. First $\psi_1(p)$ = $2pp' \leq \phi(p)+\frac{1}{2}$. Next we use Lemma 4.5 and (4.1) to prove that

(4.19)
$$
\psi_{k+1} \leq \phi(p) + \frac{1}{\sqrt{k+1}} \left(\frac{1}{2} + \sum_{i=1}^{k+1} \frac{K}{i \sqrt{i}} \right)
$$

implies

(4.20)
$$
\varphi_{k+2}(p) \leq \phi(p) + \frac{1}{\sqrt{k+2}} \left(\frac{1}{2} + \sum_{i=1}^{k+2} \frac{K}{i \sqrt{i}} \right);
$$

 K is a constant satisfying (4.11).

Since the function on the right hand side of (4.20) is concave, (4.20) implies

(4.21)
$$
\psi_{k+2}(p) = \text{Cav }\varphi_{k+2}(p) \leq \phi(p) + \frac{1}{\sqrt{k+2}} \left(\frac{1}{2} + \sum_{i=1}^{k+2} \frac{K}{i \sqrt{i}} \right).
$$

Hence we have proved (4.19) by induction for $k = 0, 1, \dots$. Finally it follows from (4.19) that

$$
\psi_{k+1} \leq \phi(p) + \frac{\beta}{\sqrt{k+1}}; \quad k = 0, 1, \cdots,
$$

holds for some constant β . This completes the proof of the Lemma.

We are now in the position to conclude the proof of our main Theorem:

PROOF OF THEOREM 2.5. By Lemmas 4.7 and 4.8

$$
\liminf_{n\to\infty}\left\{\sup_{x_0^n\in\mathcal{M}_p^n}\left[\frac{1}{\sqrt{n}}V(\chi_0^n)\right]\right\}\geq \liminf_{n\to\infty}\tilde{\phi}_n(p)\geq \phi(p)
$$

and by Theorem 4.3 and Lemma 4.9

$$
\limsup_{n\to\infty}\left\{\sup_{\chi_0^n\in\mathcal{M}_p^n}\left[\frac{1}{\sqrt{n}}V(\chi_0^n)\right]\right\}\leq \limsup_{n\to\infty}\psi_n(p)\leq \phi(p).
$$

COROLLARY 4.10. *The n-martingales* $\chi_0^n(p)$ constructed in the proof of Lemma 4.7 *satisfy*

$$
\lim_{n\to\infty}\left[\frac{1}{\sqrt{n}}V(\chi_0^n(p))\right]=\phi(p); \qquad 0\leq p\leq 1,
$$

with the speed of convergence of the order of $1/\sqrt{n}$.

In fact by construction

$$
\frac{1}{\sqrt{n}}V(\chi_0^n(p))=\tilde{\phi}_n(p); \qquad 0\leq p\leq 1; \qquad n=1,2,\cdots
$$

while by Lemmas 4.8, 4.9, and Proposition 4.6

$$
(4.22) \t\t \t\t \phi(p) - \alpha/\sqrt{n} \le \tilde{\phi}_n(p) \le \psi_n(p) \le \phi(p) + \beta/\sqrt{n}
$$

APPENDIX

Proofs of Lemmas 4.4 and 4.5

We intend to prove in this Appendix two technical statements (namely, Lemmas 4.4 and 4.5) about the function $\phi(p)$ defined on $0 \leq p \leq 1$ by

(A.1)
$$
\phi(p) = \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}x_p^2) \text{ where } \int_{-\infty}^{x_p} \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}x^2) dx = p.
$$

We start by examining the derivatives of $\phi(p)$:

PROPOSITION A.1. *For* $0 \leq p \leq 1$ *and* $\phi(p)$ *and* x_p *defined by* (A.1)

(a) $\phi'(p) = -x_p$, (b) $x'_p = 1/\phi(p)$, (c) $\phi''(p) = -1/\phi(p) = -x'_{p}$, (d) $\phi^{(3)}(p) = -x_p/\phi^2(p)$, (e) $\phi^{(4)}(p) = -(1 + 2x_p^2)/\phi^3(p)$,

- (g) $\phi^{(6)}(p) = -(4x_p^2+7)(6x_p^2+1)/\phi^5(p),$
- (h) $\phi^{(2n)}(p) \leq 0;$ $n = 1, 2, \cdots$.

PROOF. (a) to (g) result from straightforward differentiation. (h) will follow if we prove that

(A.2)
$$
\phi^{(2n)}(p) = \frac{-1}{\phi^{2n-1}(p)} \sum_{p=0}^{n-1} a_p x_p^{2p}
$$

where $a_i \ge 0$ for $j = 1, \dots, n - 1$.

We prove (A.2) by induction on n. By (c) it is true for $n = 1$. Assume it for n, then

$$
\phi^{(2n+1)}(p) = \frac{-1}{\phi^{2n}(p)} \left[\sum_{j=0}^{n-2} \left[(2n-1)a_j + 2(j+1)a_{j+1} \right] x_p^{2j+1} + (2n-1)a_{n-1} x_p^{2n-1} \right]
$$

=
$$
\frac{-1}{\phi^{2n+1}(p)} \sum_{j=0}^{n-1} \beta_j x_p^{2j+1},
$$

where $\beta_i \geq 0$; $j = 0, \dots, n - 1$. Consequently

$$
\phi^{(2n+2)}(p) = \frac{-1}{\phi^{2n+1}(p)} \left[2n \sum_{j=0}^{n-1} \beta_j x_p^{2j+2} + \sum_{j=0}^{n-1} (2j+1) \beta_j x_p^{2j} \right]
$$

$$
= \frac{-1}{\phi^{2n+1}(p)} \sum_{j=0}^{n+1} \gamma_j x_p^{2j}
$$

where $\gamma_i \geq 0$; $j = 0, \dots, n + 1$. This concludes the proof of the proposition.

PROPOSITION A.2. *If for n = 1, 2,* \cdots *we define p_n by*

(A.3)
$$
\exp(-\frac{1}{2}x_{p_n}^2) = 1/n
$$
 and $p_n \leq \frac{1}{2}$,

then there exists n_0 *s.t. for any* $n \geq n_0$

(A.4)
$$
p_n \leq p \leq p'_n \Rightarrow \phi(p)/\sqrt{n} \leq \min(p, p').
$$

PROOF. First, by our definition of x_p ; $p = (1/\sqrt{2\pi}) \int_{-\infty}^{\infty} \exp(-\frac{1}{2}x^2)$ we have

$$
p_n \leqq p \leqq p'_n \Leftrightarrow x_p^2 \leqq x_{p_n}^2 \Leftrightarrow \exp\left(-\frac{1}{2}x_p^2\right) \geqq \exp\left(-\frac{1}{2}x_{p_n}^2\right),
$$

hence $p_n \leq p \leq p'_n \Leftrightarrow \exp(-\frac{1}{2}x_p^2) \geq 1/n$ may now be written as

$$
\exp\left(-\frac{1}{2}x_p^2\right) \geq \frac{1}{n} \Rightarrow \frac{1}{\sqrt{n}} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}x_p^2\right)
$$

$$
\leq \min\left(\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{x_p} \exp\left(-\frac{1}{2}x^2\right) dx, \frac{1}{\sqrt{2\pi}} \int_{x_p}^{\infty} \exp\left(-\frac{1}{2}x^2\right) dx\right).
$$

The statement of the right hand side is

(A.5)
$$
\frac{1}{\sqrt{n}} \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}x_p^2) \leq \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{-|x_p|} \exp(-\frac{1}{2}x^2) dx.
$$

We may therefore consider just, say, $x_p \le 0$ and prove (replacing x_p by y)

(A.6)
$$
\frac{1}{\sqrt{n}} \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}y^2) \leq \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{y} \exp(-\frac{1}{2}x^2) dx
$$

whenever $\exp(-\frac{1}{2}y^2) \leq 1/n$ and $y \leq 0$.

Now (A.6) is true (for all $n \ge 1$) whenever $-1 \le y \le 0$. This is because it is true for $y = -1$ (direct computation), and the left hand side is concave on $-1 \le y \le 0$ and has a smaller slope than the right hand side which is convex on $-1 \le y \le 0$ (see Fig. 1). For $y < -1$; d/dy [exp($-\frac{1}{2}y^2$)] is positive and increasing hence at any point $y < -1$, the part of the tangent left to y lies below the line $exp(-\frac{1}{2}x^2)$. It intersects the coordinates axis at $y + 1/y$ (Fig. 2). The integral on the right hand side of $(A.6)$ can then be bounded by

$$
\int_{-\infty}^{y} \exp(-\frac{1}{2}x^2) dx \geq -\frac{1}{2y} \exp(-\frac{1}{2}y^2).
$$

It suffices therefore to prove that

$$
\exp(-\frac{1}{2}y^2) \geq 1/n \Rightarrow -1/2y \geq 1/\sqrt{n}.
$$

In fact

$$
\exp(-\frac{1}{2}y^2) \ge 1/n \Rightarrow |y| = -y \le \sqrt{2\log n} \Rightarrow -1/2y \ge 1/2\sqrt{2\log n}
$$

and since $(\log n)/n \to 0$ let n_0 be s.t. $n \ge n_0 \Rightarrow -1/2y \ge 1/\sqrt{n}$ and we have thus proved (A.4) for $n \geq n_0$.

Fig. 1.

PROOF OF LEMMA 4.4. Using Proposition A.1 we expand the first term in the left hand side of (4.10) as follows:

$$
\frac{1}{2}(\phi(p+x)+\phi(p-x)) = \phi(p) + \frac{x^2}{2}\phi''(p) + \frac{x^4}{4!} \cdot \frac{1}{2}(\phi^{(4)}(p+\delta) + \phi^{(4)}(p-\delta))
$$
\n(A.7)\n
$$
= \phi(p) - \frac{x^2}{2\phi(p)} - \frac{x^4}{4!} \cdot \frac{1}{2} \left(\frac{1+2x_{p+6}^2}{\phi^3(p+\delta)} + \frac{1+2x_{p-8}^2}{\phi^3(p-\delta)} \right)
$$

where $0 \leq \delta \leq x$.

Clearly it is enough to prove (4.10) for $n \ge n_0$ for fixed n_0 and then modify the constant c to make (4.10) true for all n .

Define p_n by $exp(-\frac{1}{2}x_{p_n}^2) = 1/n$ and $p_n \leq \frac{1}{2}$ then by Proposition A.2, $x =$ $\phi(p)/\sqrt{n}$ is in the domain of maximization in (4.10) for $n \ge n_0$, hence denoting the left hand side of (4.10) by A, we get by using (A.7) for $x = \phi(p)/\sqrt{n}$

$$
A \geq \frac{1}{\sqrt{n+1}} \left\{ \frac{\phi(p)}{\sqrt{n}} + \sqrt{n} \left[\phi(p) - \frac{\phi^2(p)}{2n\phi(p)} \right] - \frac{1}{2} \left(\frac{1 + 2x_{p+8}^2}{\phi^2(p+8)} + \frac{1 + 2x_{p-8}^2}{\phi^2(p-8)} \right) \frac{\phi^4(p)}{4! \, n^2} \right\},
$$

$$
A - \phi(p) \geq \phi(p) \left[\frac{\sqrt{n+1}/2\sqrt{n} - \sqrt{n+1}}{\sqrt{n+1}} \right]
$$

$$
- \frac{1}{2} \left(\frac{1 + 2x_{p+8}^2}{\phi^3(p+8)} + \frac{1 + 2x_{p-8}^2}{\phi^3(p-8)} \right) \frac{\phi^4(p)}{4! \, n^2 \sqrt{n+1}}.
$$

Since the first term on the right hand side is positive, the second is negative and $\sqrt{n+1} \geq 1$ we get

(A.8)
$$
A - \phi(p) \geq -\frac{1}{2} \left(\frac{1 + 2x_{p+\delta}^2}{\phi^3(p+\delta)} + \frac{1 + 2x_{p-\delta}^2}{\phi^3(p-\delta)} \right) \frac{\phi^4(p)}{4! n^2}
$$

where $0 \le \delta \le \phi(p)/\sqrt{n}$.

Notice now that $\phi^{(4)}(p) = -(1 + 2x_p^2)/\phi^3(p)$ is a negative and concave function (since $\phi^{(6)}(p) \le 0$). Also since $\phi(p)$ is symmetric around $p = \frac{1}{2}(\phi(p) = \phi(p'))$ and since $x_p = -x_p$, $\phi^{(4)}$ is also symmetric around $p = \frac{1}{2}$. It follows that for $p \leq \frac{1}{2}$

$$
\frac{1}{2}(\phi^{(4)}(p-\delta)+\phi^{(4)}(p+\delta))\geq \phi^{4}(p-\delta),
$$

and by $(A.8)$

$$
A - \phi(p) \geq -\frac{1 + 2x_{p-\delta}^2}{\phi^3(p-\delta)} \cdot \frac{\phi^4(p)}{4! \, n^2}
$$

which is

(A.9)
$$
A - \phi(p) \ge -(1 + 2x_{p-8}^2) \frac{1}{\sqrt{2\pi}} \exp(\frac{3}{2}x_{p-8}^2 - 2x_p^2) \frac{1}{4! n^2}.
$$

Now $1 + 2x^2 \leq 8 \exp(x^2/4)$ for $-\infty < x < \infty$ hence

(A.10)
$$
A - \phi(p) \ge \frac{8}{\sqrt{2\pi}4! \, n^2} \exp(\frac{7}{4}x_{p-8}^2 - 2x_p^2).
$$

We now establish the existence of a constant \tilde{K} such that $x_p - x_{p-\delta} \leq \tilde{K}/\sqrt{n}$ holds for $p_n \leq p \leq p'_n$, $p \leq \frac{1}{2}$ and *n* sufficiently large. Since $0 \leq \delta \leq \phi(p)/\sqrt{n}$ and since $x_p - x_{p-\delta}$ is monotonically increasing with δ we have to show that $\Delta \leq \tilde{K}/\sqrt{n}$ where $\Delta = x_p - x_{p-q(p)/\sqrt{n}}$. Letting $y = x_p \leq 0$ we claim in other words that $\phi(p)/\sqrt{n} = \int_{x-\Delta}^{y} (1/\sqrt{2\pi}) \exp(-\frac{1}{2}x^2) dx$ implies $\Delta \leq \tilde{K}/\sqrt{n}$.

In fact for $-1 \le y \le 0$ we have

$$
1/\sqrt{2\pi n}\geq \phi(p)/\sqrt{n}=\int_{y-\Delta}^{y}\frac{1}{\sqrt{2\pi}}\exp\left(-\frac{1}{2}x^2\right)dx\geq \frac{1}{\sqrt{2\pi e}}\Delta,
$$

which implies $\Delta \leq \sqrt{e}/\sqrt{n}$.

For $y \le -1$ the tangent to $(1/\sqrt{2\pi})\exp(-\frac{1}{2}x^2)$ at $x = y$ lies below the function and intersects the x axis at $y + 1/y$ (see Fig. 3) forming a triangular area $(-1/2y)(1/\sqrt{2\pi})exp(-\frac{1}{2}y^2) = \phi(p)/2|y|$.

Now $p \geq p_m$ implies $|y| = |x_p| \leq |x_p| = \sqrt{2 \log n} \leq \frac{1}{2} \sqrt{n}$ for *n* sufficiently large, hence the triangular area is $\geq \phi(p)/\sqrt{n}$ which implies $\Delta \leq 1/|y|$. The area of the shaded trapezoid is $\phi(p) (2 - |y|) \Delta/2$, therefore

$$
\frac{\phi(p)}{\sqrt{n}} = \frac{1}{\sqrt{2\pi}} \int_{y-\Delta}^{y} \exp\left(-\frac{1}{2}x^2\right) dx \geq \phi(p) (2-|y|\Delta) \frac{\Delta}{2} \geq \phi(p) \cdot \frac{\Delta}{2}.
$$

Fig. 3.

This completes the proof of $\Delta \leq \tilde{K}/\sqrt{n}$ for a suitable constant \tilde{K} and n sufficiently large. From here we get

$$
x_{p-\delta}^2=(x_p-\Delta)^2\leqq \left(x_p-\frac{\tilde{K}}{\sqrt{n}}\right)^2=x_p^2-\frac{2\tilde{K}x_p}{\sqrt{n}}+\frac{\tilde{K}^2}{n},
$$

and

(A.12)
$$
\frac{7}{4}x_{p-8}^2 - 2x_p^2 = -\frac{1}{4}x_p^2 - 7Kx_p/2\sqrt{n} + 7K^2/4n.
$$

Since $x_p < 0$, the right hand side has a maximum (with respect to n) at n_0 , hence

(A.13)
$$
\frac{7}{4}x_{p-6}^2 - 2x_p^2 \leq -\frac{1}{4}x_p^2 - 7Kx_p/2\sqrt{n_0} + 7K^2/4n_0 \leq K
$$

where K is the maximum of the parabola (in x_p) at the right hand side. Combining (A.13) and (A.10) we finally obtain the existence of a constant $C_1 > 0$ s.t.

$$
(A.14) \tA \geq \phi(p) - C_1/2 \tfor n \geq n_0 \tand p_n \leq p \leq p'_n.
$$

It remains to establish (A.14) also for $p \leq p_n$ or $p \geq p'_n$. In this case, by the definition of p_n : $\exp(-\frac{1}{2}x_p^2) \leq 1/n$ and therefore $\phi(p) \leq 1/n \sqrt{2\pi}$. So (choosing $x = 0$

$$
\frac{1}{\sqrt{n+1}} \max_{\substack{0 \le x \le p \\ 0 \le x \le p'}} [\sqrt{n}\frac{1}{2}(\phi(p+x)+\phi(p-x))+x] \ge \frac{\sqrt{n}}{\sqrt{n+1}}\phi(p)
$$

$$
\ge \phi(p)-\phi(p)(1-\sqrt{n}/\sqrt{n+1}) \ge \phi(p)-\phi(p)/(2n+1)
$$

$$
\ge \phi(p)-1/n(2n+1)\sqrt{2\pi} \ge \phi(p)-C_2/n^2
$$

for some constant C_2 .

Choose now C_3 s.t. $A \ge \phi(p) - C_3/n^2$ for $1 \le n \le n_0$ and finally choose $c = \max(C_1, C_2, C_3)$. This completes the proof of Lemma 4.4.

PROOF OF LEMMA 4.5. We have to prove the existence of a constant $K > 0$ s.t. for $0 \leq p \leq 1$

$$
(A.15) \frac{1}{\sqrt{n+1}} \max_{(\xi,\eta)\in S(p)} \left[\sqrt{n} \frac{\eta}{\xi+\eta} \phi(p+\xi) + \sqrt{n} \frac{\xi}{\xi+\eta} \phi(p-\eta) + \frac{2\xi\eta}{\xi+\eta} \right] \leq \phi(p) + K/n^2
$$

where $S(p) = \{(\xi, \eta) | 0 \le \xi \le p' ; 0 \le \eta \le p\}.$

Since ϕ is continuous and *S(p)* is compact, the maximum in (A.15) is achieved, say, at $({\xi_0}, {\eta_0})$. From $(d\phi/dp)_{1-} = -\infty$; $(d\phi/dp)_{0+} = +\infty$ it follows that $\xi_0 \neq p'$ and $\eta_0 \neq p$. Furthermore we claim that if $pp' \neq 0$ then $\xi_0 \neq 0$ and $\eta_0 \neq 0$. In fact, denote the function to be maximized in (A.15) by $F(\xi, \eta)$, then $F(0, \eta) = F(\xi, 0) = \sqrt{n} \phi(p)$ while

$$
\max_{(\xi,\eta)\in S(p)} F(\xi,\eta) \geq \max_{0\leq x\leq p\wedge p'} F(x,x) = \sqrt{n}\phi(p) + \max_{0\leq x\leq p\wedge p'} [O(x^2) + x]) > \sqrt{n}\phi(p).
$$

We conclude that (ξ_0, η_0) is a local maximum of $F(\xi, \eta)$ in $S(p)$. Equating first partial derivatives to 0 yields

$$
-\frac{\eta_0}{(\xi_0 + \eta_0)^2} \sqrt{n} \phi(p + \xi_0) + \frac{\eta_0}{\xi_0 + \eta_0} \sqrt{n} \phi'(p + \xi_0)
$$
\n(A.16)\n
$$
+\frac{\eta_0}{(\xi_0 + y_0)^2} \sqrt{n} \phi(p - \eta_0) + \frac{2\eta_0^2}{(\xi_0 + \eta_0)^2} = 0,
$$
\n
$$
\frac{\xi_0}{(\xi_0 + \eta_0)^2} \sqrt{n} \phi(p + \xi) - \frac{\xi_0}{(\xi_0 + \eta_0)^2} \sqrt{n} \phi(p - \eta)
$$
\n(A.17)\n
$$
-\frac{\xi_0}{\xi_0 + \eta_0} \sqrt{n} \phi'(p - \eta_0) + \frac{2\xi_0^2}{(\xi_0 + \eta_0)^2} = 0.
$$

Dividing (A.16) by $\eta_0/(\xi_0 + \eta_0)$, (A.17) by $\xi_0/(\xi_0 + \eta_0)$ and adding the results we get $\sqrt{n}[\phi'(p + \xi_0) - \phi'(p - \eta_0)] + 2 = 0.$

Recalling that $\phi'(p) = -x_p$ we have

(A.18)
$$
x_{p+\xi_0}-x_{p-\eta_0}=2/\sqrt{n}.
$$

By the mean value theorem

 $x_{p+\xi_0} - x_{p-\eta_0} = [(p+\xi_0)-(p-\eta_0)]x'_{(\theta(p+\xi_0)+(1-\theta)(p-\eta_0))}$

for some $0 \le \theta \le 1$.

Using (A.18) and recalling that $x'(p) = 1/\phi(p)$ we get

(A.19)
$$
\xi_0 + \eta_0 = (2/\sqrt{n})\phi[\theta(p + \xi_0) + (1 - \theta)(p - \eta_0)].
$$

Now

$$
\frac{\phi[\theta(p+\xi_0)+(1-\theta)(p-\eta_0)]}{\phi(p)} = \exp\{-\frac{1}{2}[x_p^2 - x_{p+\theta\xi_0-(1-\theta)\eta_0}^2]\}
$$
\n(A.20)\n
$$
= \exp\{-\frac{1}{2}[x_p + x_{p+\theta\xi_0-(1-\theta)\eta_0}][x_p - x_{p+\theta\xi_0-(1-\theta)\eta_0}]\}.
$$

Since x_p is monotonically increasing in p we get from $(A.20)$

$$
[x_p + x_{p+\theta\xi_0 - (1-\theta)\eta_0}][x_p - x_{p+\theta\xi_0 - (1-\theta)\eta_0}] \leq (2/\sqrt{n})(2|x_p| + 2/\sqrt{n})
$$

and by (A.19) and (A.20) therefore

(a.21) ~:,, + *rio <=* (2/x/n) *4)(p)exp(21xp I/x/n)"* exp(2/n).

Denote

$$
G(\xi,\eta)=\frac{\eta}{\xi+\eta}\,\phi(p+\xi)+\frac{\xi}{\xi+\eta}\,\phi(p-\eta).
$$

Expanding $\phi(p + \xi)$ and $\phi(p - \eta)$ yields the following expansion for $G(\xi, \eta)$:

$$
G(\xi, \eta) = \phi(p) + \frac{1}{2}\xi\eta\phi''(p) + \frac{1}{6}\xi\eta(\xi - \eta)\phi'''(p) + \frac{1}{24}\xi\eta(\xi^2 - \xi\eta + \eta^2)\phi^{(4)}(p) + \frac{1}{120}\left[\frac{\eta\xi^5}{\xi + \eta}\phi^{(5)}(p + \sigma_1\xi) - \frac{\xi\eta^5}{\xi\eta}\phi^{(5)}(p - \sigma_2\eta)\right],
$$

where $0 \le \sigma_1 \le 1$ and $0 \le \sigma_2 \le 1$.

First consider the last term in (A.22) which we denote by $K(p; \xi, \eta)$. Since $\phi^{(5)}$ is decreasing we have by Proposition A.1

(A.23)
$$
K(p;\xi,\eta) \leq -\frac{1}{120}\xi\eta(\xi^2+\eta^2)(\xi-\eta)x_p(7+6x_p^2)/\phi^4(p).
$$

By (A.21), since max $({\xi \eta}, {\xi^2 + \eta^2}) \leq ({\xi + \eta})^2$ and ${\xi - \eta \leq \xi + \eta}$ we have

$$
|K(p; \xi_0, \eta_0)| \leq \tfrac{1}{120} \left[\frac{2}{\sqrt{n}} \phi(p) \exp(2|x_p|/\sqrt{n}) + \frac{2}{\sqrt{n}} \right]^5 x_p (7 + 6x_p^2)/\phi^4(p)
$$

$$
(A.24) = \frac{4}{15\sqrt{n^5}} [\phi(p)x_p(7+6x_p^2) \exp(10|x_p|/\sqrt{n})+10/\sqrt{n}]
$$

=
$$
\frac{4}{15\sqrt{2\pi n^5}} [x_p(7+6x_p^2) \exp(10|x_p|/\sqrt{n}+10/\sqrt{n}-\frac{1}{2}x_p^2)].
$$

The last expression is clearly a bounded function of x_p , hence

$$
(A.25) \qquad |K(p;\xi_0,\eta_0)| \leq K_1/n^2,
$$

for some constant K_1 .

By (A.22) and (A.25), using Proposition A.1 we obtain

$$
(A.26) \qquad G(\xi_0, \eta_0) \leq \phi(p) - \xi_0 \eta_0/2 \phi(p) - \frac{1}{6} \xi_0 \eta_0 (\xi_0 - \eta_0) x_p / \phi^2(p) -
$$

$$
- \frac{1}{24} \xi_0 \eta_0 (\xi_0^2 - \xi_0 \eta_0 + \eta_0^2) (1 + 2 x_p^2) / \phi^3(p) + K_1/n^2.
$$

Therefore

$$
(A.27) \qquad \max_{(\xi,\eta)\in S(p)} \left[G(\xi,\eta) + \frac{1}{\sqrt{n}\xi+\eta} \right] \leq \phi(p) + K_1/n^2 + \max_{(\xi,\eta)\in S(p)} D(\xi,\eta)
$$

where

$$
D(\xi,\eta)=\frac{1}{\sqrt{n}\,\xi+\eta}-\frac{\xi\eta}{2\phi(p)}-\frac{1}{6}\xi\eta(\xi-\eta)\frac{x_p}{\phi^2(p)}-\frac{1}{24}\xi\eta(\xi^2-\xi\eta+\eta^2)\frac{1+2x_p^2}{\phi^3(p)}
$$

Observe that $D(0, \eta) = D(\xi, 0) = 0$; $D(\varepsilon, \varepsilon) > 0$ for $\varepsilon > 0$ sufficiently small. Also $D(\xi, \eta) \rightarrow -\infty$ as $\xi \rightarrow \infty$ or $\eta \rightarrow \infty$. It follows that D restricted to the nonnegative orthant has a global maximum which is also a local maximum. Equating first derivatives of $D(\xi, \eta)$ to 0 gives

(A.29)
$$
0 = \frac{\partial D}{\partial \xi} = \frac{1}{\sqrt{n}} \frac{2\eta^2}{(\xi + \eta)^2} - \frac{\eta}{2\phi(p)} - \frac{1}{6} [\eta(\xi - \eta) + \xi \eta] \frac{x_p}{\phi^2(p)} - \frac{\eta}{24} (3\xi^2 - 2\xi\eta + \eta^2) \frac{1 + 2x_p^2}{\phi^3(p)},
$$

(A.30)
$$
0 = \frac{\partial D}{\partial \eta} = \frac{1}{\sqrt{n}} \frac{2\xi^2}{(\xi + \eta)^2} - \frac{\xi}{2\phi(p)} + \frac{1}{6} [\xi(\eta - \xi) + \xi\eta] \frac{x_p}{\phi^2(p)} - \frac{\xi}{24} (3\eta^2 - 2\xi\eta + \xi^2) \frac{1 + 2x_p^2}{\phi^3(p)}.
$$

Adding

$$
\frac{1}{\eta} \frac{\partial D}{\partial \xi} + \frac{1}{\xi} \frac{\partial D}{\partial \eta}
$$

we get

$$
\textbf{(A.31)} \ 2/\sqrt{n}(\xi+\eta)-1/\phi(p)-\frac{1}{2}(\xi-\eta)x_p/\phi^2(p)-\frac{1}{6}(\xi^2-\xi\eta+\eta^2) \times (1+2x_p^2)/\phi^3(p)=0.
$$

Subtracting

$$
\frac{1}{\eta} \frac{\partial D}{\partial \xi} - \frac{1}{\xi} \frac{\partial D}{\partial \eta}
$$

we get

(A.32)
$$
2(\eta - \xi)/\sqrt{n}(\xi + \eta)^2 - \frac{1}{6}(\xi + \eta)x_p/\phi^2(p) - \frac{1}{12}(\xi^2 - \eta^2)(1 + 2x_p^2)/\phi^3(p) = 0.
$$

By dividing (A.32) by $(\xi^2 - \eta^2)$ and eliminating $(\eta - \xi)$ we obtain

(A.33)
$$
\eta - \xi = \frac{x_p/\phi^2(p)}{\frac{12}{\sqrt{n}(\xi + \eta)^3} + (\frac{1}{2} + x_p^2)/\phi^3(p)}.
$$

Replacing in (A.31) $(\xi^2 - \xi \eta + \eta^2)$ by $\frac{1}{4}(\xi + \eta)^2 + \frac{3}{4}(\xi - \eta)^2$, and $(\xi - \eta)$ by its value according to (A.33), we get

$$
\frac{2}{\sqrt{n}(\xi+\eta)} - \frac{1}{\phi(p)} + \frac{x_p^2/\phi^4(p)}{24/\sqrt{n}(\xi+\eta)^3 + (1+2x_p^2)/\phi^3(p)} - \frac{1}{24} \frac{1+2x_p^2}{\phi^3(p)} (\xi+\eta)^2 - \frac{1}{8} \frac{1+2x_p^2}{\phi^3(p)} \left[\frac{x_p/\phi^2(p)}{12/\sqrt{n}(\xi+\eta)^3 + (\frac{1}{2}+x_p^2)/\phi^3(p)} \right]^2 = 0.
$$

It is easily verified that the expression in (A.34) tends to $+\infty$ as $({\xi + \eta}) \rightarrow 0$. On the other hand, the last two terms are always negative and the third is bounded

from above by $[1/\phi(p)]$ {Max_x $[x/(1+2x^2)]$ } which is $1/2\sqrt{2}\phi(p)$. So if we denote the left hand side of (A.34) by $L(\xi, \eta)$ we can assert that

(A.35)
$$
L(\xi, \eta) \leq 2/\sqrt{(\xi + \eta)} - 1/\phi(p) + 1/2\sqrt{2}\phi(p).
$$

The right hand side of (A.35) is non-negative if and only if

$$
\xi + \eta \leq \alpha \phi(p)/\sqrt{n} \quad \text{where} \quad \alpha = 2/(1 - 1/2\sqrt{2}) \sim 3.1.
$$

It follows therefore from (A.35) and (A.34) that any solution (ξ, η) of (A.31) and (A.33) must satisfy

(A.36) ~c + n =< a~b (p)/~/n.

By (A.33) we get that at the maximum (ξ, η)

$$
|\eta - \xi| \leq \frac{|x_p|/\phi^2(p)}{12n/\phi^3(p)\alpha^3 + (\frac{1}{2} + x_p^2)/\phi^3(p)}
$$

=
$$
\frac{|x_p|\phi(p)}{12n/\alpha^3 + (\frac{1}{2} + x_p^2)} < \frac{|x_p|\phi(p)\alpha^3}{12n}.
$$

Being interested in obtaining an upper bound for the global maximum of D we replace its last two terms by an upper bound at the maximum. The resulting function will have a maximum which is greater than or equal to that of D. Now the last term of D (in (A.28)) is not positive and as for the third term, by (A.36) and (A.37),

$$
\frac{1}{6}\xi\eta(\xi-\eta)\frac{x_p}{\phi^2(p)} \leq \frac{1}{6}\alpha^2 \frac{\phi^2(p)|x_p|\phi(p)\alpha^3}{n-12n\phi^2(p)}
$$

$$
= \frac{\alpha^5}{72}\frac{1}{n^2}\text{Max}\frac{1}{\sqrt{2\pi}}|x|\exp(-\frac{1}{2}x^2) \leq K_2/n^2.
$$

We conclude that

(A.38)
$$
\max_{\xi, \eta} D(\xi, \eta) \leq K_2/n^2 + \max_{\xi, \eta} D_1(\xi, \eta)
$$

where K_2 is a constant and

(A.39)
$$
D_1(\xi,\eta)=\frac{2\xi\eta}{\xi+\eta}\cdot\frac{1}{\sqrt{n}}-\frac{\xi\eta}{2\phi(p)}.
$$

Equating $\partial D_1/\partial \xi$ and $\partial D_1/\partial \eta$ to 0 we get

$$
\frac{1}{\eta} \frac{\partial D_1}{\partial \xi} = \frac{2\eta}{(\xi + \eta)^2} \cdot \frac{1}{\sqrt{n}} - \frac{1}{2\phi} = 0,
$$

$$
\frac{1}{\xi} \frac{\partial D_1}{\partial \eta} = \frac{2\xi}{(\xi + \eta)^2} \cdot \frac{1}{\sqrt{n}} - \frac{1}{2\phi} = 0
$$

which imply $\xi = \eta = \phi/\sqrt{n}$ and hence

(A.40)
$$
\max_{(\xi,\eta)\in S(p)} D_1(\xi,\eta) \leq \phi(p)/n - \phi(p)/2n = \phi(p)/2n.
$$

By (A.27), (A.38) and (A.40)

(A.41)
$$
\max_{(\xi,\eta)\in S(p)}\bigg[G(\xi,\eta)+\frac{1}{\sqrt{n}}\frac{2\xi\eta}{\xi+\eta}\bigg]\leq \phi(p)(1+1/2n)+(K_1+K_2)/n^2.
$$

Coming back to (A.15) we have now by (A.41)

$$
\frac{1}{\sqrt{n+1}}\max_{(\xi,\eta)\in S(p)}\left[\sqrt{n}\frac{\eta}{\xi+\eta}\phi(p+\xi)+\sqrt{n}\frac{\xi}{\xi+\eta}\phi(p-\eta)+\frac{2\xi\eta}{\xi+\eta}\right]
$$
\n(A.42)
\n
$$
=\frac{\sqrt{n}}{\sqrt{n+1}}\max_{(\xi,\eta)\in S(p)}\left[G(\xi,\eta)+\frac{1}{\sqrt{n}}\cdot\frac{2\xi\eta}{\xi+\eta}\right]\leq \phi(p)\left(1+\frac{1}{2n}\right)\frac{\sqrt{n}}{\sqrt{n+1}}+\frac{K_1+K_2}{n^2}.
$$

Now notice that $\sqrt{1 + 1/n} \ge 1 + 1/2n - 1/8n^2$, therefore

$$
\left(1+\frac{1}{2n}\right)\frac{\sqrt{n}}{\sqrt{n+1}}-1=\frac{1+1/2n}{\sqrt{1+1/n}}-1\leq \frac{1/8n^2}{1+1/2n-1/8n^2}\leq K_3/n^2
$$

where K_3 is a constant. It follows that

$$
(A.43) \ \ \phi(p)\Big(1+\frac{1}{2n}\Big)\sqrt{n}/\sqrt{n+1}+K_1/n^2 \leq \phi(p)+\phi(p)K_2/n^2+K_1/n^2
$$

$$
\leq \phi(p)+K/n^2
$$

where K is a constant. (A.42) and (A.43) conclude the proof of Lemma 4.5,

REFERENCES

1. J. L. Doob, *Stochastic Processes,* John Wiley and Sons Inc., New York, 1953.

2. W. Feller, *An Introduction to Probability Theory and its Applications I* (third edition), John Wdey and Sons Inc., New York, 1968.

3. J. F. Mertens and S. Zamir, The *normal distribution and repeated games,* to appear in Internat. J. Game Theory (1977).

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